

PROBLEM 1. (a)

$$m_i = \mathbb{E} \left[ \sum_j b_j Y_j \middle| H = i \right] = \sum_j b_j \mathbb{E} [Y_j | H = i] = s_i \sum_j a_j b_j \mathbb{E} [Z_j] = s_i \sum_j a_j b_j, \quad (1)$$

$$v_i = \text{Var} \left[ \sum_j b_j Y_j \middle| H = i \right] \stackrel{(\star)}{=} \sum_j b_j^2 \text{Var} [Y_j | H = i] = s_i^2 \sum_j a_j^2 b_j^2 \text{Var} [Z_j] = s_i^2 \sum_j a_j^2 b_j^2, \quad (2)$$

where in  $(\star)$  we use the fact that the  $Y_j$  are uncorrelated. Combining eqs. (1) and (2) gives

$$q = \frac{(m_0 - m_1)^2}{v_0 + v_1} = \frac{(s_0 - s_1)^2}{s_0^2 + s_1^2} \frac{\left( \sum_j a_j b_j \right)^2}{\sum_j a_j^2 b_j^2}. \quad (3)$$

(b) Let  $\mathbf{c} = [a_1 b_1, \dots, a_n b_n] \in \mathbb{R}^n$ .

$$\left( \sum_j a_j b_j \right)^2 = \left( \sum_j c_j \right)^2 = |\langle \mathbf{c}, \mathbf{1} \rangle|^2 \stackrel{(\star)}{\leq} \|\mathbf{c}\|_2^2 \|\mathbf{1}\|_2^2 = n \sum_j a_j^2 b_j^2,$$

where  $(\star)$  is due to the Cauchy-Schwarz inequality. Plugging the above into eq. (3) gives

$$q \leq n \frac{(s_0 - s_1)^2}{(s_0^2 + s_1^2)}. \quad (4)$$

(c) Quality  $q$  is maximized when eq. (4) holds with equality, which is true if and only if  $\mathbf{c}$  and  $\mathbf{1}$  are colinear, i.e.,

$$b_j = \frac{\lambda}{a_j}, \quad (5)$$

for some constant  $\lambda \in \mathbb{R}$ .

(d)

$$f_{Y|H}(y|i) = \prod_{j=1}^n f_{Y_j|H}(y_j|i) = \prod_{j=1}^n f_{Z_j} \left( \frac{y_j}{s_i a_j} \right) = \exp \left( -\frac{1}{s_i} \sum_{j=1}^n \frac{y_j}{a_j} \right) \stackrel{eq. (5)}{=} \exp \left( -\frac{T}{\lambda s_i} \right).$$

As the joint distribution only depends on  $Y = [Y_1, \dots, Y_n]$  through  $T$ , the statistic is sufficient.

PROBLEM 2. (a) A simple orthonormal basis for  $\mathcal{W} = \{w_0, w_1, w_2, w_3\}$  is

$$\mathcal{B} = \{\psi_0(t) = \psi(t), \psi_1(t) = \psi(t-1)\}, \quad \psi(t) = \mathbb{1}_{[0,1]}(t).$$

Using  $\mathcal{B}$ , the MAP optimal decision rule is given by

$$\hat{H} = \arg \max_{i \in \{0,1,2,3\}} \langle Y, c_i \rangle - \frac{1}{2} \|c_i\|_2^2,$$

where  $Y = [\langle R, \psi_0 \rangle, \langle R, \psi_1 \rangle] \in \mathbb{R}^2$  and  $c_i = [\langle w_i, \psi_0 \rangle, \langle w_i, \psi_1 \rangle] \in \mathbb{R}^2$ . In practice we cannot obtain  $Y$  as above since we don't have a matched filter for  $\psi$ . Notice however that  $\psi(t) = h_0(t) + h_1(t)$  such that

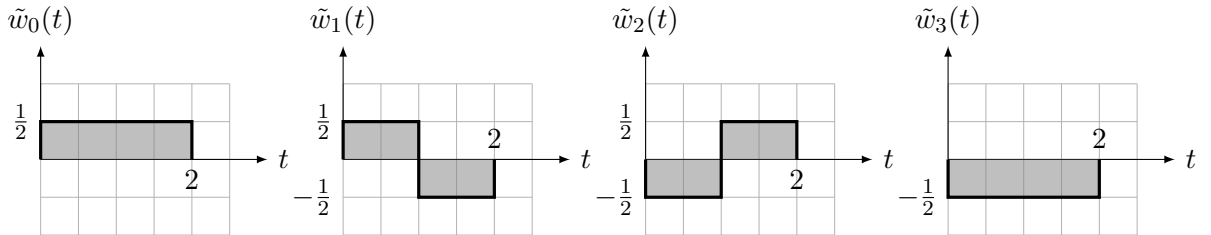
$$\begin{aligned} \langle R, \psi_0 \rangle &= \int R(t) \psi_0(t) dt = \int R(t) h_0(t) dt + \int R(t) h_1(t) dt \\ &= (R * h_1)(1) + (R * h_0)(1), \\ \langle R, \psi_1 \rangle &= \int R(t) \psi_0(t-1) dt = \int R(t) h_0(t-1) dt + \int R(t) h_1(t-1) dt \\ &= (R * h_1)(2) + (R * h_0)(2). \end{aligned}$$

Therefore the optimal MAP decoder is still achievable by choosing  $Y = [Y_{10} + Y_{00}, Y_{11} + Y_{01}]$  with  $t_{10} = t_{00} = 1$  and  $t_{11} = t_{01} = 2$ .

(b) The minimum energy signal set  $\tilde{\mathcal{W}} = \{\tilde{w}_0, \tilde{w}_1, \tilde{w}_2, \tilde{w}_3\}$  is obtained by subtracting the mean signal  $m(t)$  from the  $w_i$ :

$$m = \frac{1}{4} \sum_i w_i = \frac{1}{2} \psi_0 + \frac{1}{2} \psi_1$$

$$\begin{aligned} \tilde{w}_0 &= w_0 - m = \frac{1}{2} \psi_0 + \frac{1}{2} \psi_1, & \tilde{w}_1 &= w_1 - m = \frac{1}{2} \psi_0 - \frac{1}{2} \psi_1, \\ \tilde{w}_2 &= w_2 - m = -\frac{1}{2} \psi_0 + \frac{1}{2} \psi_1, & \tilde{w}_3 &= w_3 - m = -\frac{1}{2} \psi_0 - \frac{1}{2} \psi_1. \end{aligned}$$



(c) Using basis  $\mathcal{B}$  given above, the codewords associated to  $\tilde{\mathcal{W}}$  are

$$\tilde{c}_0 = \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix}, \quad \tilde{c}_1 = \begin{bmatrix} \frac{1}{2} \\ -\frac{1}{2} \end{bmatrix}, \quad \tilde{c}_2 = \begin{bmatrix} -\frac{1}{2} \\ \frac{1}{2} \end{bmatrix}, \quad \tilde{c}_3 = \begin{bmatrix} -\frac{1}{2} \\ -\frac{1}{2} \end{bmatrix}.$$

Recognizing a QAM constellation with minimum codeword distance  $d = 1$ , the error probability is given by

$$P_e(\tilde{\mathcal{W}}) = 2Q\left(\frac{1}{\sqrt{2N_0}}\right) - Q\left(\frac{1}{\sqrt{2N_0}}\right)^2.$$

- (d)  $\mathcal{W}$  and  $\widetilde{\mathcal{W}}$  are linked through an isometric transform, hence their error probabilities are identical:

$$P_e(\mathcal{W}) = P_e(\widetilde{\mathcal{W}}) = 2Q\left(\frac{1}{\sqrt{2N_0}}\right) - Q\left(\frac{1}{\sqrt{2N_0}}\right)^2.$$

PROBLEM 3. (a) By the Cauchy-Schwarz inequality:

$$\langle w_i, w_j \rangle^2 \leq \|w_i\|_2^2 \|w_j\|_2^2. \quad (6)$$

For  $i \neq j$ , eq. (6) simplifies to  $\beta^2 \leq 1$ , thus proving the claim.

- (b) Based on the hint, we have

$$\begin{aligned} \|w_0 + w_1 + w_2\|_2^2 &= \|w_0\|_2^2 + \|w_1 + w_2\|_2^2 + 2\langle w_0, w_1 + w_2 \rangle \\ &= \|w_0\|_2^2 + \|w_1\|_2^2 + \|w_2\|_2^2 + 2\langle w_0, w_1 \rangle + 2\langle w_0, w_2 \rangle + 2\langle w_1, w_2 \rangle \\ &= 3 + 6\beta. \end{aligned} \quad (7)$$

By the non-negativity property of  $\|\cdot\|_2$ , we must have  $3+6\beta \geq 0$ , which implies  $\beta \geq -\frac{1}{2}$ .

- (c) The minimum energy signal set  $\widetilde{\mathcal{W}} = \{\tilde{w}_0, \tilde{w}_1, \tilde{w}_2\}$  is obtained by subtracting the mean signal  $m(t)$  from the  $w_i$ :

$$m = \frac{1}{3} \sum_i w_i.$$

$$\begin{aligned} \tilde{w}_0 &= w_0 - m = +\frac{2}{3}w_0 - \frac{1}{3}w_1 - \frac{1}{3}w_2, \\ \tilde{w}_1 &= w_1 - m = -\frac{1}{3}w_0 + \frac{2}{3}w_1 - \frac{1}{3}w_2, \\ \tilde{w}_2 &= w_2 - m = -\frac{1}{3}w_0 - \frac{1}{3}w_1 + \frac{2}{3}w_2. \end{aligned}$$

Let  $\{w_i, w_j, w_k\}$  and  $\{\tilde{w}_i, \tilde{w}_j, \tilde{w}_k\}$  be some arbitrary relabeling of  $\{w_0, w_1, w_2\}$  and  $\{\tilde{w}_0, \tilde{w}_1, \tilde{w}_2\}$  respectively:

$$\begin{aligned} \|\tilde{w}_i\|_2^2 &= \|w_i - m\|_2^2 = \|w_i\|_2^2 + \|m\|_2^2 - 2\langle w_i, m \rangle \\ &= \|w_i\|_2^2 + \left\| \frac{1}{3}w_i + \frac{1}{3}w_j + \frac{1}{3}w_k \right\|_2^2 - 2\langle w_i, \frac{1}{3}w_i + \frac{1}{3}w_j + \frac{1}{3}w_k \rangle \\ &\stackrel{eq. (7)}{=} \|w_i\|_2^2 + \frac{1}{9}(3 + 6\beta) - \frac{2}{3}(\|w_i\|_2^2 + \langle w_i, w_j \rangle + \langle w_i, w_k \rangle) \\ &= 1 + \frac{1}{9}(3 + 6\beta) - \frac{2}{3}(1 + 2\beta) = \frac{2}{3}(1 - \beta) = E, \end{aligned} \quad (8)$$

$$\begin{aligned} \langle \tilde{w}_i, \tilde{w}_j \rangle &= \left\langle \frac{2}{3}w_i - \frac{1}{3}w_j - \frac{1}{3}w_k, -\frac{1}{3}w_i + \frac{2}{3}w_j - \frac{1}{3}w_k \right\rangle \\ &= -\frac{2}{9}\|w_i\|_2^2 + \frac{5}{9}\langle w_i, w_j \rangle - \frac{1}{9}\langle w_i, w_k \rangle - \frac{2}{9}\|w_j\|_2^2 - \frac{1}{9}\langle w_j, w_k \rangle + \frac{1}{9}\|w_k\|_2^2 \\ &= -\frac{1}{3}(1 - \beta) \stackrel{eq. (8)}{=} -\frac{E}{2}. \end{aligned} \quad (9)$$

- (d)  $\widetilde{\mathcal{W}}$  is related to  $\mathcal{W} = \{w_0, w_1, w_2\}$  through an isometric transform, hence they have the same error probability  $P_e$ . We will therefore consider  $\widetilde{\mathcal{W}}$  below.

Notice from item c that  $\|\tilde{w}_i\|_2^2 = E$  and  $\frac{\langle \tilde{w}_i, \tilde{w}_j \rangle}{\|\tilde{w}_i\|_2 \|\tilde{w}_j\|_2} = -\frac{1}{2} = \cos(120^\circ)$ , therefore  $\tilde{\mathcal{W}}$  is indeed a 3-PSK constellation and we can quantify its error rate  $P_e$  using  $e_3(\cdot)$ . To this end, let us define an orthonormal basis  $\mathcal{B} = \{\psi_0, \psi_1\}$  for  $\tilde{\mathcal{W}}$ :

$$\psi_0 = \frac{\tilde{w}_0}{\|\tilde{w}_0\|_2} = \frac{\tilde{w}_0}{\sqrt{E}}, \quad \psi_1 = \frac{\tilde{w}_1 - \langle \tilde{w}_1, \psi_0 \rangle \psi_0}{\|\tilde{w}_1 - \langle \tilde{w}_1, \psi_0 \rangle \psi_0\|_2} = \frac{\sqrt{E}}{2} \tilde{w}_0 + \tilde{w}_1.$$

Using  $\mathcal{B}$ , the sufficient statistic  $Y \in \mathbb{R}^2$  behaves as

$$Y|H=0 \sim \mathcal{N}\left(\begin{bmatrix} \sqrt{E} \\ 0 \end{bmatrix}, \frac{N_0}{2} I_2\right).$$

Equivalently  $\tilde{Y} = \sqrt{\frac{2}{N_0}} Y$  behaves as

$$\tilde{Y}|H=0 \sim \mathcal{N}([A, 0], I_2),$$

with  $A = \sqrt{2E/N_0}$ . Therefore communicating with  $\mathcal{W}$  over an AWGN channel of power spectral density  $\frac{N_0}{2}$  gives rise to an error rate

$$P_e = e_3\left(\sqrt{\frac{2E}{N_0}}\right) = e_3\left(\sqrt{\frac{4(1-\beta)}{3N_0}}\right).$$

PROBLEM 4. (a) Conditioned on  $H$  and  $t_0$ ,  $Y$  follows the Gaussian distribution

$$Y|H, t_0 \sim \mathcal{N}\left((w_H * h)(t_0), \frac{N_0}{2} \|h\|_2^2\right), \quad (10)$$

$$(w_H * h)(t_0) = \int_{t_0 - \frac{1}{2}}^{t_0 + \frac{1}{2}} w_H(t) dt = \frac{(-1)^H}{2}. \quad (11)$$

Given the above, we have

$$P_e|H, t_0 = Q\left(\frac{|(w_H * h)(t_0)|}{\sqrt{N_0/2} \|h\|_2}\right) = Q\left(\sqrt{\frac{1}{2N_0}}\right),$$

$$P_e = \mathbb{E}_{H, t_0} [P_e|H, t_0] = Q\left(\sqrt{\frac{1}{2N_0}}\right).$$

(b) eq. (10) still holds here, but now  $\|h\|_2 = \sqrt{2}$  and  $\mathbb{E}[Y|H, t_0]$  is given by

$$(w_H * h)(t_0) = \int_{t_0 - 1}^{t_0 + 1} w_H(t) dt = (-1)^H.$$

The error probability therefore becomes

$$P_e|H, t_0 = Q\left(\frac{|(w_H * h)(t_0)|}{\sqrt{N_0/2} \|h\|_2}\right) = Q\left(\sqrt{\frac{1}{N_0}}\right),$$

$$P_e = \mathbb{E}_{H, t_0} [P_e|H, t_0] = Q\left(\sqrt{\frac{1}{N_0}}\right).$$

(c) eq. (10) still holds here, but we will simplify terms differently:

$$\|h\|_2^2 = \int_{-\infty}^0 h^2(t) dt + \int_0^{\infty} h^2(t) dt = E_- + E_+,$$

$$\begin{aligned} (w_H * h)(0.5) &= \int w_H(t) h\left(\frac{1}{2} - t\right) dt = (-1)^H \int_{-\frac{1}{2}}^{\frac{1}{2}} h\left(\frac{1}{2} - t\right) dt \\ &= (-1)^H \int_0^1 h(\alpha) d\alpha = (-1)^H A_+, \end{aligned}$$

$$\begin{aligned} (w_H * h)(-0.5) &= \int w_H(t) h\left(-\frac{1}{2} - t\right) dt = (-1)^H \int_{-\frac{1}{2}}^{\frac{1}{2}} h\left(-\frac{1}{2} - t\right) dt \\ &= (-1)^H \int_{-1}^0 h(\alpha) d\alpha = (-1)^H A_-. \end{aligned}$$

The error probability therefore becomes

$$P_e|H, t_0 = Q\left(\frac{|(w_H * h)(t_0)|}{\sqrt{N_0/2} \|h\|_2}\right) = \begin{cases} Q\left(\sqrt{\frac{2A_+^2}{N_0(E_- + E_+)}}\right), & t_0 = 0.5 \\ Q\left(\sqrt{\frac{2A_-^2}{N_0(E_- + E_+)}}\right), & t_0 = -0.5 \end{cases},$$

$$P_e = \mathbb{E}_{H, t_0} [P_e|H, t_0] = \frac{1}{2} Q\left(\sqrt{\frac{2A_+^2}{N_0(E_- + E_+)}}\right) + \frac{1}{2} Q\left(\sqrt{\frac{2A_-^2}{N_0(E_- + E_+)}}\right).$$

(d)

$$A_+^2 = \left(\int_0^1 h(t) \cdot 1 dt\right)^2 \leq \int_0^1 h^2(t) dt \int_0^1 1 dt \leq \int_0^{\infty} h^2(t) dt = E_+,$$

where the first inequality is due to the Cauchy-Schwarz theorem. A similar argument applied to  $(A_-, E_-)$  also implies  $A_-^2 \leq E_-$ .

(e)

$$\begin{aligned} P_e &\stackrel{(c)}{=} \frac{1}{2} Q\left(\sqrt{\frac{A_+^2}{\sigma^2(E_- + E_+)}}\right) + \frac{1}{2} Q\left(\sqrt{\frac{A_-^2}{\sigma^2(E_- + E_+)}}\right) \geq Q\left(\sqrt{\frac{A_-^2 + A_+^2}{2\sigma^2(E_- + E_+)}}\right) \\ &\stackrel{(d)}{\geq} Q\left(\sqrt{\frac{1}{2\sigma^2}}\right) = Q\left(\sqrt{\frac{1}{N_0}}\right). \end{aligned}$$

As the lower bound is achieved by the filter used in (b), the former is optimal.